

# Final Review Sheet Answers

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The proofs/calculations of most exercises here are omitted. Again, refer to your notes if unsure; all of the theoretical results have been discussed in lecture notes or in the textbook.

## 1 Some common aspects of densities

1. *Constant and variable parts of densities:*

$$X \sim \text{Gamma}(4, 6); \int_0^\infty 5x^3 e^{-6x} dx = 5\left(\frac{\Gamma(4)}{6^4}\right)$$

2. *Independence and dependence:*

In the continuous case, we may show that  $f_X(x, y) = f_X(x)f_Y(y)$ , **and** that the support of  $X$  and  $Y$  do not depend on each other (i.e., for all  $(x, y)$ , we have that  $f_{X,Y}(x, y) > 0$  if and only if  $f_X(x) > 0$  and  $f_Y(y) > 0$ .)

An easy way to show that two variables are dependent is to find  $x, y$  such that  $P(Y = y) > 0$  but  $P(Y = y|X = x) = 0$ . The idea is to find a value of  $X$  that makes a particular value of  $Y$  impossible.

3. *Working with densities:*

(a) The CDF of  $X$  is given by  $F_X(x) = \int_{-\infty}^x f_X(t)dt$ . By the fundamental theorem of calculus, it follows that  $f_X(x) = F'_X(x)$ .

(b)  $f_X(x) = \int_{-\infty}^{\infty} f_{X,Y}(x, y)dy$

## 2 Relationships Between Distributions

1.  $cX$ , where  $X \sim \text{Exp}(\lambda)$ ,  $c > 0$ :

$\text{Exp}\left(\frac{\lambda}{c}\right)$ . The methods we have used include using the CDF of  $X$ , using the change of variable method, and using the MGF of  $X$  and properties of MGFs.

2. Consider  $X = U_{(3)}$  and  $Y = U_{(7)}$ , the 3rd and 7th order statistics of 10 iid Unif(0, 1) RVs. What is the distribution of  $X/Y$ ?

We proved in lecture that for this type of example (ratio of joint Uniform (0,1) order statistics, where the numerator is the smaller one), the resulting distribution is a Beta. In particular, this example follows the Beta(3,4) distribution.

3.  $X^2 + Y^2$ , where  $X, Y$  are independent standard Normal. What is  $\sqrt{X^2 + Y^2}$ ?

$\text{Exp}\left(\frac{1}{2}\right)$ ; standard Rayleigh

### 3 Symmetry

1. Under some conditions, we can quickly recognize the expectation of a random variable  $X$  to be zero. What are they?

The distribution/density of  $X$  must be symmetric about the origin, and  $\mathbb{E}(|X|) < \infty$ ; i.e. the expectation must be defined.

2. Let  $X, Y$  be independent  $\mathcal{N}(0, \sigma^2)$  random variables. Without using  $\Phi$ , find  $P(X + 2Y > 0, X > 0)$ . What is the reason behind this answer?

Using the rotational symmetry of the joint distribution of  $(X, Y)$ ,

$$P(X + 2Y > 0, X > 0) = \frac{\frac{\pi}{2} + \arctan(\frac{1}{2})}{2\pi}$$

### 4 Simplifying an infinite sum

1. Let  $Y$  have density  $f_Y(y) = \frac{\lambda}{2}e^{-\lambda|y|}$ , for  $y \in \mathbb{R}$ . With little computation, find  $\mathbb{E}(|Y|)$  and  $\mathbb{E}(Y)$ .

If we look at the graph of the density of  $Y$ , or through the change-of-variable formula, we observe that  $|Y| \sim \text{Exp}(\lambda)$ . Thus  $\mathbb{E}(|Y|) = \frac{1}{\lambda}$ , and  $\mathbb{E}(Y) = 0$  by symmetry.

### 5 Could You Rephrase That?

1. Let  $X \sim \text{Gamma}(r, \lambda)$ , where  $r$  is an integer. Use what we know about the Poisson process to obtain the CDF of  $X$ .

$$\begin{aligned} P(X < t) &= P(N_t \geq r) \\ &= 1 - P(N_t < r) \\ &= 1 - \sum_{k=0}^{r-1} e^{-\lambda t} \frac{(\lambda t)^k}{k!} \end{aligned}$$

2. Let  $X$  be an arbitrary random variable with an invertible CDF  $F_X$ . What is the random variable formed by  $F_X(X)$ ?

Let  $Z = F_X(X)$ . Then,

$$\begin{aligned} F_Z(z) &= P(Z \leq z) \\ &= P(F_X(X) \leq z) \\ &= P(F_X^{-1}(F_X(X)) \leq F_X^{-1}(z)) \\ &= P(X \leq F_X^{-1}(z)) \\ &= F_X(F_X^{-1}(z)) \\ &= z, \quad z \in [0, 1] \end{aligned}$$

We conclude that  $Z \sim \text{Unif}(0,1)$ .

## 6 Some Useful Results

1. Let  $X \sim \text{Exp}(\lambda_X)$ ,  $Y \sim \text{Exp}(\lambda_Y)$ . What is  $P(X < Y)$ ?  $\frac{\lambda_X}{\lambda_X + \lambda_Y}$
2. Find the distribution of  $\min\{X_1, X_2, \dots, X_n\}$ , where the  $X_i$ 's are independent  $\text{Exp}(\lambda)$  variables.  
Let  $M$  denote the minimum. Then  $M \sim \text{Exp}(n\lambda)$ . Note this result generalizes to the case where the rates are all different; you simply add the rate parameters.
3. Suppose cars and trucks arrive at a bridge according to independent Poisson processes with rates  $\lambda_c$  and  $\lambda_r$  per minute respectively. Given that  $n$  vehicles arrive in  $t$  minutes, what is the distribution of  $N_{c,t}$ , the number of cars to arrive by time  $t$ ?

The easiest way to proceed here is using the conditional probability rule. We find that  $N_{c,t} | N_t = n \sim \text{Binom}(n, \frac{\lambda_c}{\lambda_c + \lambda_r})$ .