

Stat 134    Lec 41

Warmup 11:00-11:10

Let  $X, Y$  std bivariate normal,  $\rho > 0$

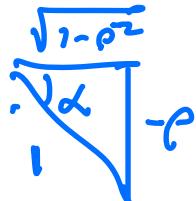
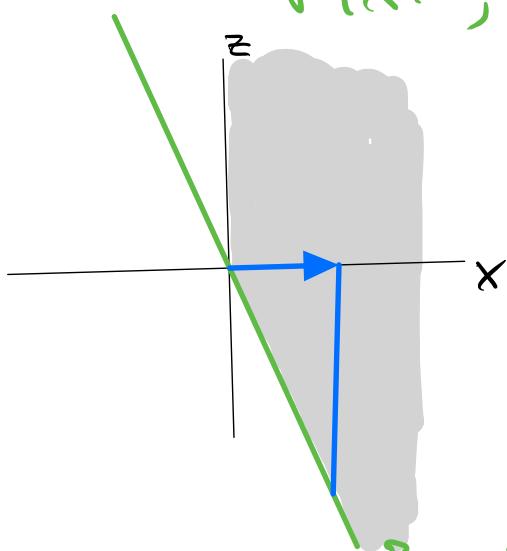
Find  $P(X > 0, Y > 0)$

Hint

$$P(X > 0, Y > 0) = P\left(X > 0, \rho X + \sqrt{1-\rho^2} Z > 0\right)$$

$$= P\left(X > 0, Z > -\frac{\rho}{\sqrt{1-\rho^2}} X\right)$$

note  $X$  and  $Z$  are uncorrelated  
so  $\text{joint}(X, Z)$  is a symmetric bell over  $X, Z$  plane,



$$Z = -\frac{\rho}{\sqrt{1-\rho^2}} X$$

$$\tan \alpha = \frac{-\rho}{\sqrt{1-\rho^2}} \Rightarrow \alpha = \tan^{-1}\left(\frac{-\rho}{\sqrt{1-\rho^2}}\right)$$

$$P\left(X > 0, Z > -\frac{\rho}{\sqrt{1-\rho^2}} X\right) = \boxed{\frac{90 + |\alpha|}{360}}$$

## Last time

- RRR week schedule  
 M Review (post questions on b-courses),  
 W OFF  
 F Review (post questions on b-courses)

} OFF  
after  
class  
on  
zoom

### Sec 6.5. Bivariate Normal

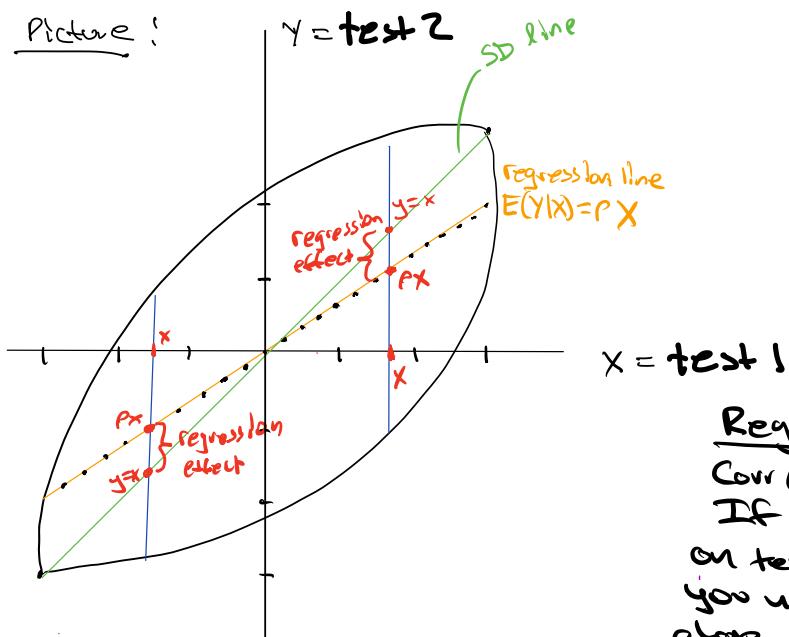
Defn (Standard Bivariate Normal Distribution)

let  $X, Z \sim N(0,1)$ ,  $-1 \leq \rho \leq 1$

$$Y = \rho X + \sqrt{1-\rho^2} Z \sim N(0,1)$$

$$\text{Cov}(X, Y) = \rho$$

## regression effect



Regression effect,  
 $\text{Cov}(\text{test 1}, \text{test 2}) = .6$   
 If 1 SD above mean  
 on test 1 then on average  
 you will be less than 1 SD  
 above average on test 2.  
 (regression line is less steep  
 than SD line).

Today (1) MGF of bivariate normal

(2) sec 6.5 Properties of bivariate normal

(3) sec 6.4 Total Variance Decomposition

## ① MGF of bivariate normal

The single and multivariate MGF is defined as:

$$M_y(t) = E(e^{tY}) \quad \text{single variable MGF}$$

$$M_{(x,y)}(s,t) = E(e^{sx+ty}) \quad \text{multivariate MGF}$$

Show that  $M_{(x,y)}(s,t) = M_x(s) M_y(t)$  iff  
 $x, y$  are independent,

$$\begin{aligned} M_{(x,y)}(s,t) &= E(e^{sx+ty}) \\ &= E(e^{sx} \cdot e^{ty}) \\ &\stackrel{\text{iff } x, y \text{ indep.}}{=} E(e^{sx}) E(e^{ty}) \\ &= M_x(s) \cdot M_y(t) \end{aligned}$$

$$E(AB) = E(A)E(B)$$

iff  $A, B$   
are indep.

Have  
 $A = e^{sx}$   
 $B = e^{ty}$

Thm Let  $(X, Y)$  be standard bivariate normal.

The MGF of  $(X, Y)$  is

$$M_{(X,Y)}(s,t) = e^{\frac{s^2}{2} + \frac{t^2}{2} + st\rho}$$

$$\begin{aligned} \text{Pf/ } M_{X,Y}(s,t) &= E[e^{sx+ty}] \\ &= E[e^{sx+t(\rho X + \sqrt{1-\rho^2}Z)}] \\ &= E[e^{(s+t\rho)x} \cdot e^{t\sqrt{1-\rho^2}z}] \\ &\stackrel{\text{independence}}{=} E[e^{(s+t\rho)x}] E[e^{t\sqrt{1-\rho^2}z}] \\ &= M_X(s+t\rho) \cdot M_Z(t\sqrt{1-\rho^2}) \\ &= e^{\frac{(s+t\rho)^2}{2}} \cdot e^{\frac{t^2\sqrt{1-\rho^2}}{2}} \end{aligned}$$

Recall that  $X \sim N(0,1)$  so  $M_X(a) = e^{\frac{a^2}{2}}$

$$\begin{aligned} &= e^{\frac{(s+t\rho)^2}{2}} \cdot e^{\frac{(t\sqrt{1-\rho^2})^2}{2}} \\ &= e^{\frac{s^2}{2} + st\rho + \frac{t^2\rho^2}{2}} \cdot e^{\frac{t^2 - t^2\rho^2}{2}} \\ &= e^{\boxed{\frac{s^2}{2} + \frac{t^2}{2} + st\rho}} \end{aligned}$$

Finish proof.

## ② Properties of Bivariate Normal

Recall that if 2 RVs  $X, Y$  are independent,  $\text{Cov}(X, Y) = 0$   
 $\Rightarrow \text{Corr}(X, Y) = 0$

However the converse is not true in general. ( $\text{Corr}(X, Y) = 0 \not\Rightarrow X, Y \text{ indep.}$ )

Let  $X \sim N(0, 1)$

$$Y = X^2$$

$X$  and  $Y$  are dependent

we show  $X, Y$  are uncorrelated,

$$\text{Cov}(X, Y) = E(XY) - E(X)E(Y)$$

$$= E(X^3) - E(X)E(X^2)$$

### Real world example

$$x = \text{height}$$

$$y = (\text{height})^2$$

If you make a scatter plot of height vs  $(\text{height})^2$  it will look uncorrelated even though height and  $(\text{height})^2$  are dependent

Find  $E(X^3)$

$$= \int_{-\infty}^{\infty} x^3 \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx$$

$\int_{-\infty}^{\infty} x^3 \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx = 0$  since  $\frac{x^3}{\sqrt{2\pi}} e^{-\frac{x^2}{2}}$  is an odd function

OR

$$M_X(t) = e^{t^2/2} \quad M_X'''(0) = t e^{t^2/2} + 2t e^{t^2/2} + t^3 e^{t^2/2} \Big|_{t=0} = 0$$

$$\Rightarrow E(X^3) = 0$$

$$\Rightarrow \text{Corr}(X, Y) = 0$$

D

Thm If  $(X, Y)$  is <sup>std</sup> bivariate normal then

$\rho = \text{Corr}(X, Y) = 0$  iff  $X, Y$  are independent.

Pf/ From the warm up we know  
for any joint distribution  $(X, Y)$ ,

$$M_{(X,Y)}(s,t) = M_X(s) M_Y(t) \text{ iff } X, Y \text{ are independent.}$$

Since  $(X, Y)$  is <sup>std</sup> bivariate normal,

$$M_{(X,Y)}(s,t) = e^{\frac{s^2}{2} + \frac{t^2}{2} + st\rho}$$

$$\begin{aligned} X, Y \text{ indep} \quad M_{(X,Y)}(s,t) &= M_X(s) M_Y(t) \Rightarrow e^{st\rho} = 1 \\ &\quad " " \quad " " \quad " " \quad " " \Rightarrow (\rho = 0) \\ &\quad " " \quad " " \quad " " \quad " " \end{aligned}$$

$$\begin{aligned} \text{Conversely, if } \rho = 0 \quad e^{\frac{s^2}{2} + \frac{t^2}{2} + st\rho} &= e^{\frac{s^2}{2} + \frac{t^2}{2}} = e^{\frac{s^2}{2}} \cdot e^{\frac{t^2}{2}} \Rightarrow X, Y \text{ indep.} \\ M_{(X,Y)}(s,t) &= M_X(s) M_Y(t) \end{aligned}$$

Recall from Lec 30 that the sum of independent normal random variables is normal.

What about dependent normal random variables?

recall,

$$M_Y(t) = E(e^{tY}) = M_{tY}(1)$$

single variable MGF

$$M_{(X,Y)}(s,t) = E(e^{sX+tY}) = M_{sX+tY}(1)$$

multivariate MGF

recall,

$$Z \sim N(\mu, \sigma^2) \text{ iff } M_Z(\omega) = e^{\mu\omega} e^{\frac{\sigma^2 \omega^2}{2}}$$

Thm Let  $X, Y \sim N(0, 1)$  and  $\text{Corr}(X, Y) = \rho$ .

$(X, Y)$  is std bivariate normal iff

$sX + tY$  is normal for all constants  $s, t$ .

Pf

$\Rightarrow$ : Suppose  $X, Y$  std bivariate normal.  
(i.e.  $Y = \rho X + \sqrt{1-\rho^2} Z$  for  $X, Z \sim N(0, 1)$ )

$$\begin{aligned}sX + tY &= sX + t(\rho X + \sqrt{1-\rho^2} Z) \\ &= (s + t\rho)X + t\sqrt{1-\rho^2} Z\end{aligned}$$

is normal since  $X, Z$  are indep normals

$\Leftarrow$ : Suppose  $sX + tY$  is normal,

Note  $E(sX + tY) = sE(X) + tE(Y) = 0$

$$\begin{aligned} \text{Var}(sX + tY) &= \text{Var}(sX) + \text{Var}(tY) + 2\text{Cov}(sX, tY) \\ &= s^2 + t^2 + 2st\rho \end{aligned}$$

$$2st\text{Cov}(X, Y)$$

$$\text{Cov}(X, Y) = \rho$$

Hence,  $sX + tY \sim N(0, s^2 + t^2 + 2st\rho)$

Recall,

$$\text{For } Z \sim N(\mu, \sigma^2) \Rightarrow M_Z^{(k)} = e^{\mu k} e^{\sigma^2 \frac{k^2}{2}}$$

$$\begin{aligned} \text{Then } M_{(X,Y)}^{(s,t)} &= M_{sX+tY}^{(1)} = e^{(s^2 + t^2 + 2st\rho)\frac{1}{2}} \\ &\quad \text{normal} \qquad \leftarrow \text{MF of } (X, Y) \\ &= e^{\frac{s^2}{2} + \frac{t^2}{2} + st\rho} \end{aligned}$$

$\Rightarrow (X, Y)$  is standard bivariate normal  $\square$

We don't need to restrict ourselves to  $X, Y \sim N(0, 1)$

Corollary — proof at end of lecture

$$\text{Let } U \sim N(\mu_U, \sigma_U^2)$$

$$V \sim N(\mu_V, \sigma_V^2)$$

$$\text{and } \text{Corr}(U, V) = \rho$$

$$(U, V) \sim BV(\mu_U, \mu_V, \sigma_U^2, \sigma_V^2, \rho) \text{ if}$$

$sU + tV$  is normal for all constants  $s, t$ .

Ex

Let  $M$  be a student's score on the mid-term at a class and  $F$  the student's score on the final at the same class.

Suppose  $(M, F) \sim \text{BV}(70, 65, 8, 10^2, 0.6)$ .

$$\begin{matrix} M & M_F & \alpha_M & \alpha_F^2 & \text{Cov}_F \\ \uparrow & \uparrow & \nearrow & \nearrow & \end{matrix}$$

Find the chance that the student scores higher on the final than the mid-term.

Hint  $P(F > M) = P(F - M > 0)$

What distribution is  $F - M$ ?  $\xrightarrow{\text{Normal}}$   
 (Since  $(M, F) \sim \text{Bivariate Normal}$ )

$$E(F - M) = 65 - 70 = -5$$

$$\text{Var}(F - M) = \text{Var}(F) + \text{Var}(M) - 2\text{Cov}(F, M)$$

$$\stackrel{!!}{=} \text{Corr}(F, M) \text{SD}(M) \text{SD}(F)$$

$$= 10^2 + 8^2 - 2(0.6)(8)(10) = 68$$

$$\text{SD}(F - M) = \sqrt{68}$$

$$\Rightarrow F - M \sim N(-5, 68)$$

$$\Rightarrow \frac{0 - E(F - M)}{\text{SD}(F - M)} = \frac{0 - (-5)}{\sqrt{68}} = \boxed{.61}$$



$$P(F > M) = P(F - M > 0) = 1 - \Phi(.61) = \boxed{.27}$$

## Appendix

### Corollary

Let  $U \sim N(\mu_U, \sigma_U^2)$

$V \sim N(\mu_V, \sigma_V^2)$

and  $\text{Corr}(U, V) = \rho$

$(U, V) \sim BV(\mu_U, \mu_V, \sigma_U^2, \sigma_V^2, \rho)$  iff

$sU + tV$  is normal for all constants  $s, t$ .

Pf/ let  $X = \frac{U - \mu_U}{\sigma_U}$  ( $U = \mu_U + X\sigma_U$ )

$Y = \frac{V - \mu_V}{\sigma_V}$  ( $V = \mu_V + Y\sigma_V$ )

$(U, V) \sim BV(\mu_U, \mu_V, \sigma_U^2, \sigma_V^2, \rho)$

iff  $(X, Y) \sim BV(0, 0, 1, 1, \rho)$

$$\begin{aligned} sU + tV &= s(\mu_U + X\sigma_U) + t(\mu_V + Y\sigma_V) \\ &= \underset{a}{s\sigma_U}X + \underset{b}{t\sigma_V}Y + \underset{\text{constant}}{s\mu_U + t\mu_V} \end{aligned}$$

this is normal for all constants  $s, t$

if  $\alpha X + \beta Y$  is normal for all constants  $\alpha, \beta$ ,

□

