

Stat 134 lec 16

Warmup 10:00-10:10

A tub of blueberry muffin batter has $\lambda = 2$ bb/in³ intensity of bb,

bb muffin 1 is 3 in³

bb muffin 2 is 4 in³

a) on average how many bb is in muffin 1?

$$\mu_1 = \lambda \cdot 3 = \boxed{6 \text{ bb}}$$

b) Find $P(5 \text{ bb in each muffin})$

$$X_1 \sim \text{Pois}(6)$$

$$X_2 \sim \text{Pois}(8)$$

$$P(X_1=5, X_2=5) = \frac{e^{-6} 6^5}{5!} \cdot \frac{e^{-8} 8^5}{5!}$$

c) Find $P(10 \text{ bb total in both muffins together})$,

$$X_1 + X_2 \sim \text{Pois}(14)$$

$$P(X_1 + X_2 = 10) = \frac{e^{-14} 14^{10}}{10!}$$

Announcements

For Monday's review, write down questions in discussion board on b-course by Sunday 8pm.

Last time

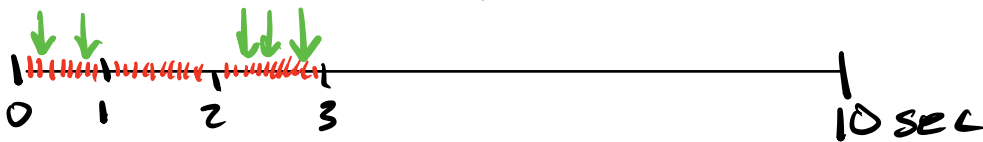
sec 3.5 Poisson distribution

$$X \sim \text{Pois}(\mu)$$

$$P(X=k) = \frac{e^{-\mu} \mu^k}{k!}$$

$$E(X) = \text{Var}(X) = \mu.$$

Poisson Process or Poisson Random Scatter (PRS):
e.g. radioactive decay of Americium 241 in 10 seconds



Assumptions

- ① no two particles arrive at the same time.
(this allows us to divide 10 sec into n small time intervals each with at most one arrival.)
- ② X is a sum of n ← large iid Bernoulli(p) trials, n small

$\mu = n p$ is avg # of arrivals in 10 sec.
 $\lambda = \mu / 10$ is the arrival rate per second.

$X =$ # arrivals in 10 seconds.

Suppose $\lambda = 4$ arrivals/sec

then $\mu = \lambda \cdot 10 = 40 \Rightarrow X \sim \text{Pois}(40)$

Americium has a long half life.

$Y =$ # arrivals in 12070 sec.

$Y \sim \text{Pois}(\lambda \cdot 12070)$
 $\lambda = 4$

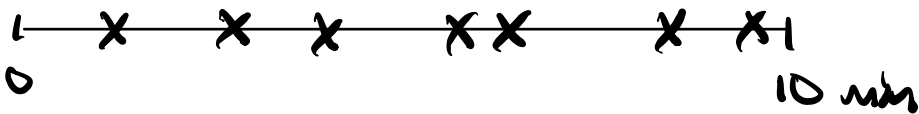
Today (1) Finish section 3.5 Poisson Thinning

(2) mid-term review

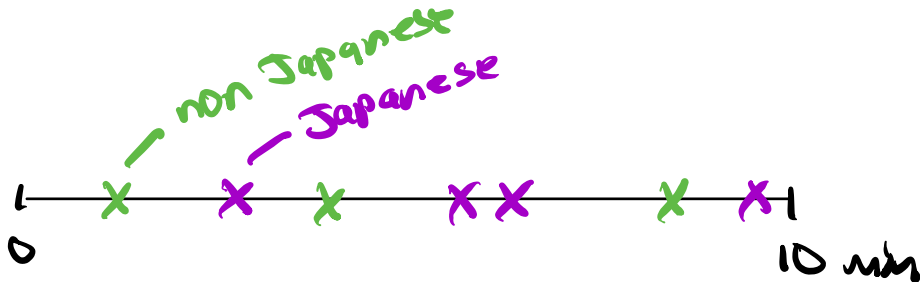
Sec 3.5 Poisson thinning

Cars arrive at a toll booth according to a Poisson process at a rate $\lambda = 3$ arrivals/min

$X = \#$ cars arriving at a toll booth in 10 min. $X \sim \text{Pois}(\lambda \cdot 10)$
 $\lambda \cdot 10 = 30$



Of cars arriving, it is known, over the long term, that 60% are Japanese imports,



Call Japanese cars a success and non Japanese a failure.

$$\# \text{ cars} \sim \text{Pois}(\lambda \cdot 10) = \text{Pois}(30)$$

$$\# \text{ Japanese imports} \sim \text{Pois}(p \lambda \cdot 10) = \text{Pois}(18)$$

$$\# \text{ non Japanese} \sim \text{Pois}(q \lambda \cdot 10) = \text{Pois}(12)$$

Ex What is the probability that in a given 10 min interval, 15 cars arrive at the booth and 10 are Japanese imports?

$$X = \# \text{ cars in 10 min} \sim \text{Pois}(30)$$

indep

$$J = \# \text{ Japanese cars in 10 min} \sim \text{Pois}(18)$$

$$nJ = \# \text{ non-Japanese} \sim \text{Pois}(12)$$

$$P(X=15, J=10) = P(nJ=5, J=10)$$

$$= P(nJ=5)P(J=10)$$

$$= \frac{e^{-12} 5^5}{5!} \cdot \frac{e^{-18} 10^{10}}{10!}$$



② midterm review

Which distributions are (approximately) a sum of a fixed number of independent Bernoulli trials?

Discrete

name and range	$P(k) = P(X = k)$ for $k \in \text{range}$	mean	variance
uniform on $\{a, a+1, \dots, b\}$ $\{1, 2, \dots, n\}$	$\frac{1}{b-a+1}$	$\frac{a+b}{2}$ $\frac{n+1}{2}$	$\frac{(b-a+1)^2 - 1}{12}$ $\frac{n^2 - 1}{12}$
Bernoulli (p) on $\{0, 1\}$	$P(1) = p; P(0) = 1 - p$	p	$p(1-p)$
binomial (n, p) on $\{0, 1, \dots, n\}$	$\binom{n}{k} p^k (1-p)^{n-k}$	np	$np(1-p)$
Poisson (μ) on $\{0, 1, 2, \dots\}$	$\frac{e^{-\mu} \mu^k}{k!}$	μ	μ
hypergeometric (n, N, G) on $\{0, \dots, n\}$ if $n \leq N$	$\frac{\binom{G}{k} \binom{N-G}{n-k}}{\binom{N}{n}}$	$\frac{nG}{N}$	$n \frac{\binom{G}{N} \binom{N-G}{N} \binom{N-n}{N-1}}$
geometric (p) on $\{1, 2, 3, \dots\}$	$(1-p)^{k-1} p$	$\frac{1}{p}$	$\frac{1-p}{p^2}$
geometric (p) on $\{0, 1, 2, \dots\}$	$(1-p)^k p$	$\frac{1-p}{p}$	$\frac{1-p}{p^2}$
negative binomial (r, p) on $\{0, 1, 2, \dots\}$	$\binom{k+r-1}{r-1} p^r (1-p)^k$	$\frac{r(1-p)}{p}$	$\frac{r(1-p)}{p^2}$

normal

$\phi(x)$

μ

σ^2

CLT \Rightarrow normal \Rightarrow sum of iid Bernoulli(p) trials

DeMorgan's rule: $(A \cap B)^c = A^c \cup B^c$
 $\Rightarrow A \cap B = (A^c \cup B^c)^c$

$$\text{So } \boxed{P(A \cap B) = 1 - P(A^c \cup B^c)}$$

Inclusion exclusion formula:

Let A_1, A_2, A_3 be dependent RVs with

$$P(A_i) = .9 \text{ for } i=1, 2, 3.$$

Find a lower bound for $P(\bigcap_{i=1}^3 A_i)$

$$P(\bigcap_{i=1}^3 A_i) = 1 - P(\bigcup_{i=1}^3 A_i^c) \text{ by De Morgan's rule.}$$

$$P(\bigcup_{i=1}^3 A_i^c) = P(A_1^c) + P(A_2^c) + P(A_3^c)$$

$$- P(A_1^c A_2^c) - P(A_2^c A_3^c) - P(A_1^c A_3^c)$$

$$+ P(A_1^c A_2^c A_3^c)$$

$$\Rightarrow P(\bigcup_{i=1}^3 A_i^c) \leq P(A_1^c) + P(A_2^c) + P(A_3^c) = 3(.1) = .3$$

$$P(\bigcap_{i=1}^3 A_i) \geq 1 - .3 = \boxed{.7}$$

ex Conditional distribution, Poisson

8. Let X_1 and X_2 be independent random variables such that for $i = 1, 2$, the distribution of X_i is Poisson (μ_i). Let m be a fixed positive integer. Find the distribution of X_1 given that $X_1 + X_2 = m$. Recognize this distribution as one of the famous ones, and provide its name and parameters.

$$\left. \begin{array}{l} X_1 \sim \text{Pois}(\mu_1) \\ X_2 \sim \text{Pois}(\mu_2) \end{array} \right\} \text{ indep} \quad P(X_1 = k) = \frac{e^{-\mu_1} \mu_1^k}{k!}$$

$X_1 | X_1 + X_2 = m$ takes values $0, 1, 2, \dots, m$

$$\begin{aligned} P(X_1 = k | X_1 + X_2 = m) &= \frac{P(X_1 = k, X_2 = m - k)}{P(X_1 + X_2 = m)} \\ &= \frac{P(X_1 = k) P(X_2 = m - k)}{P(X_1 + X_2 = m)} \\ &= \frac{\frac{e^{-\mu_1} \mu_1^k}{k!} \cdot \frac{e^{-\mu_2} \mu_2^{m-k}}{(m-k)!}}{\frac{e^{-(\mu_1 + \mu_2)} (\mu_1 + \mu_2)^m}{m!}} \\ &= \binom{m}{k} \left(\frac{\mu_1}{\mu_1 + \mu_2} \right)^k \left(\frac{\mu_2}{\mu_1 + \mu_2} \right)^{m-k} \end{aligned}$$

$$\Rightarrow X_1 | X_1 + X_2 = m \sim \text{Bin} \left(m, \frac{\mu_1}{\mu_1 + \mu_2} \right)$$

Problem 4 (conditional probability)

Two jars each contains r red marbles and b blue marbles. A marble is chosen at random from the first jar and placed in the second jar. A marble is then randomly chosen from the second jar. Find the probability this marble is red.

X = the first marble color

Y = the second marble color.

R = red

B = blue

$$\frac{r+1}{r+b+1} \cdot \frac{r}{r+b}$$

$$P(Y=R) = P(Y=R|X=R)P(X=R) + P(Y=R|X=B)P(X=B)$$

$$\frac{r}{r+b+1} \cdot \frac{b}{r+b}$$

expectation question

An urn contains 90 marbles, of which there are 20 greens, 20 blacks and 50 red marbles. Tom draw marble without replacement until the 6th green marble. Let $X = \#$ of marbles drawn. Example: **G G G B R B G G B R G** with $x = 11$. Find $\mathbb{E}[X]$.

Hint First find the expected number of marbles until the 1st green marble.

What is the min and max of X ?

1, 71
 $P = 1/21$

$$X = I_1 + \dots + I_{70} + 1$$

$$\text{where } I_i = \begin{cases} 1 & \text{if } i^{\text{th}} \text{ nongreen} \\ & \text{is before } 1^{\text{st}} \text{ green} \\ 0 & \text{else} \end{cases}$$

2nd nongreen must go here

$$\rightarrow G_1 - G_2 - \dots - G_{20} -$$

$P = 1/21$ since there are 21 slots the 2nd nongreen can go.

$$\Rightarrow \mathbb{E}(X) = 70 \cdot (1/21) + 1$$

To solve original problem by symmetry you expect # nongreen marbles between first and second green also to be $70 \cdot (1/21) + 1$. Similar between 2nd and 3rd green etc

Hence answer is

$$\boxed{6 \cdot \left[70 \left(\frac{1}{21} \right) + 1 \right]}$$

